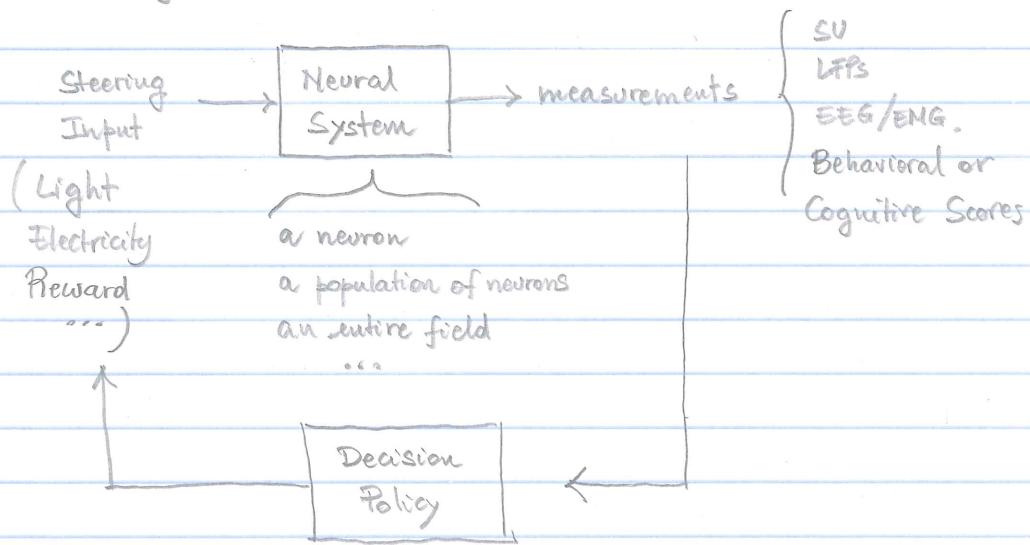


(1)

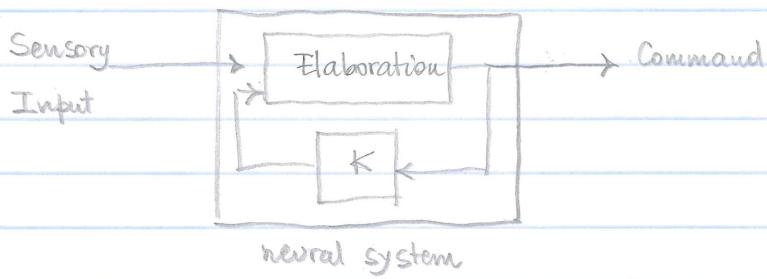
LECTURE 1

* Neural Control : what do we mean ?

1) Control = Exogenous Intervention



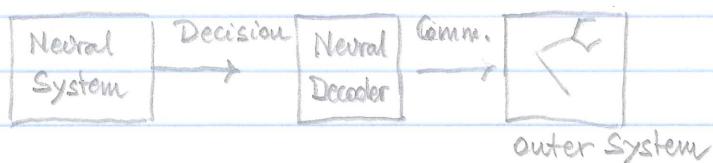
2) Control = Internal Regulation



Ex: Motor Commands

Learning, etc.

3) Control = Actuation of Outer System



(2)

In all three cases we need to provide:

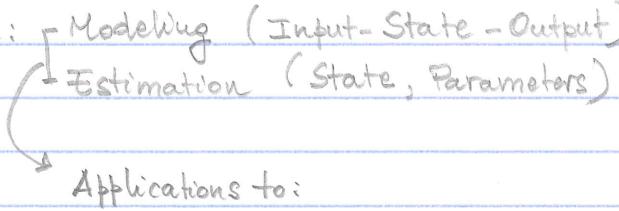
- A representation of what the neural system does in response to an input
- A characterization of the uncertainty that affects such response



We need to introduce models (to represent) and estimation tools (to characterize) - We also need objectives to constrain the models



The topics of the course are:



□

* Tools that we need to carry on the estimation

1) Least-Squares

$$Y = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix}$$

- measurements
(e.g.; firing rates
in numerous repetitions)

$$X \triangleq \begin{bmatrix} x_1 \\ x_2 \\ \vdots \\ x_m \end{bmatrix}$$

- explicative
variables
($m < n$)

For instance, $m=1$ and x_1 is the actual firing rate (which is

(3)

supposed constant) - We have:

$$\text{error}_1 = y_1 - x_1$$

$$\text{error}_2 = y_2 - x_1$$

$$\vdots$$

$$\text{error}_n = y_n - x_1$$

$\underbrace{\quad}_{E}$

$$\leftrightarrow E = Y - \begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix} x_1 = Y - A x_1$$

$\stackrel{n \text{ times}}{\overbrace{\quad \quad \quad \quad}}$
 \hat{A}

In general, if $m > 1$, then A is $n \times m$. Moreover coefficients of A can be arbitrary:

$$E = Y - AX$$

Problem: What is the estimation \hat{X} that minimizes the norm of E ?

Solution:

$$\begin{aligned} \|E\|^2 &\triangleq \sum_{i=1}^n \text{error}_i^2 = (Y - AX)^T (Y - AX) = \\ &= Y^T Y - X^T A^T Y - Y^T A X + X^T A^T A X = f(X) \quad (*) \end{aligned}$$

The min of $f(X)$ (which is a scalar function) has derivatives equal to zero \Rightarrow We impose:

$$\left. \begin{array}{l} \frac{\partial f}{\partial x_1} = 0 \\ \frac{\partial f}{\partial x_2} = 0 \\ \vdots \\ \frac{\partial f}{\partial x_m} = 0 \end{array} \right\}$$

If X were scalar, we would derive from (*):

$$-A^T Y - Y^T A + 2A^T A X = 0$$

\Updownarrow Since it's scalar

$$2A^T A X - 2A^T Y = 0$$

$$\hat{X} = (A^T A)^{-1} A^T Y$$

(4)

The same result holds if X is a $m \times 1$ vector:

$$\hat{X} = (A^T A)^{-1} A^T Y \quad (**)$$

Note: $(**)$ is a projection from the n -dim space of Y to the m -dim space of X (\Rightarrow We are linearly combining all measurements in Y to extract an information about X)



What if not all the measurements in Y are equally good?

We introduce a matrix of weights W , e.g., a diagonal matrix:

$$W \triangleq \begin{bmatrix} \sigma_1 & & & 0 \\ & \sigma_2 & & \\ & & \ddots & \\ 0 & & & \sigma_n \end{bmatrix} \quad \text{and we assign each entry to one specific error:}$$

$$\|WE\|^2 = (WY - WAX)^T (WY - WAX)$$

By repeating the argument above, we have:

$$2A^T W^T W Y - 2A^T W^T W A X = 0$$



$$\hat{X} = (A^T W^T W A)^{-1} A^T W^T W Y$$

Note: $WTW = \begin{bmatrix} \sigma_1^2 & & & 0 \\ & \sigma_2^2 & & \\ & & \ddots & \\ 0 & & & \sigma_n^2 \end{bmatrix} \Rightarrow$ We are weighting each measurement a quadratic value

Note: E is a $n \times 1$ vector $\Rightarrow R = \text{cov}(E)$ is a $n \times n$ matrix \Rightarrow It can be

(5)

shown that, by choosing $W^T W = R^{-1}$, the solution \hat{X} is the best unbiased estimator of X :

$$\hat{X} = (A^T R^{-1} A)^{-1} A^T R^{-1} Y$$

This choice of $W^T W$ is convenient also for another reason:

\hat{X} is an estimation \Rightarrow There is an error between \hat{X} and X of the actual X

\Rightarrow The covariance of this error is given by:

$$P \triangleq E \left[(X - \hat{X})(X - \hat{X})^T \right] = (A^T R^{-1} A)^{-1}$$

↑
expected
value

Example: $R \triangleq \begin{bmatrix} \sigma^2 & & & \\ & \ddots & & \\ & & \ddots & 0 \\ 0 & & & \sigma^2 \end{bmatrix} \quad \left(\begin{array}{c} \\ \text{---} \\ \text{---} \\ n \times n \end{array} \right) \Rightarrow A^T R^{-1} A = [1 \dots 1] \begin{bmatrix} \frac{1}{\sigma^2} & & & \\ & \ddots & & \\ & & \frac{1}{\sigma^2} & \\ & & & 1 \end{bmatrix} [1 \dots 1]$

$$A \triangleq \begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix} \quad n \times 1$$

$$= \sum_{i=1}^n \frac{1}{\sigma^2} = \frac{n}{\sigma^2}$$

$\Rightarrow P = (A^T R^{-1} A)^{-1} = \sigma^2/n$ - The larger n , the lower the covariance of the error

Note: The estimation of P and \hat{X} proceeds in "batch" mode (\Rightarrow All measurements must be available in Y) and assumes that the A is constant \Rightarrow What if A changes in time?

(6)

Assume that there are consecutive batches of measurements

$$\begin{array}{c}
 \begin{matrix} Y_1 & Y_2 & \dots & Y_k \end{matrix} \\
 \sim \quad \sim \quad \dots \quad \sim \\
 n_1 \times 1 \quad n_2 \times 1 \quad \dots \quad n_k \times 1 \\
 \text{model} \quad \downarrow \quad \downarrow \quad \downarrow \\
 A_1 X \quad A_2 X \quad A_k X \\
 \downarrow \\
 E = \begin{bmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_k \end{bmatrix} - \begin{bmatrix} A_1 \\ A_2 \\ \vdots \\ A_k \end{bmatrix} X \quad - \text{Block partition}
 \end{array}$$

$n \triangleq \sum_{i=1}^k n_i$
 A_1, A_2, \dots, A_k are known
 Y_1, Y_2, \dots, Y_k are independent

If we consider the first batch, we can write:

$$P_1^{-1} = A_1^T R_1^{-1} A_1 \quad \text{where } R_1 \text{ is the covariance matrix of the error computed for the first batch}$$

$$\hat{X}_1 = P_1 A_1^T R_1^{-1} Y_1$$

If we consider the first two batches, we have:

$$\begin{aligned}
 P_2^{-1} &= [A_1^T \ A_2^T] \begin{bmatrix} R_1^{-1} & 0 \\ 0 & R_2^{-1} \end{bmatrix} \begin{bmatrix} A_1 \\ A_2 \end{bmatrix} \\
 &= A_1^T R_1^{-1} A_1 + A_2^T R_2^{-1} A_2 = P_1^{-1} + A_2^T R_2^{-1} A_2 \quad (\text{a})
 \end{aligned}$$

$$\hat{X}_2 = P_2 \begin{bmatrix} A_1^T \ A_2^T \end{bmatrix} \begin{bmatrix} R_1^{-1} & 0 \\ 0 & R_2^{-1} \end{bmatrix} \begin{bmatrix} Y_1 \\ Y_2 \end{bmatrix} =$$

(7)

$$= P_2 \left(A_1^T R_1^{-1} Y_1 + A_2^T R_2^{-1} Y_2 \right)$$

$$= P_2 \left(P_1^{-1} \hat{X}_1 + A_2^T R_2^{-1} Y_2 \right)$$

$$= P_2 \left(P_1^{-1} \hat{X}_1 - A_2^T R_2^{-1} A_2 \hat{X}_1 + A_2^T R_2^{-1} Y_2 \right)$$

↑
Replace
 P_1^{-1} with (a)

$$= \hat{X}_1 + P_2 A_2^T R_2^{-1} (Y_2 - A_2 \hat{X}_1)$$

$\underbrace{\phantom{P_2 A_2^T R_2^{-1}}}_{\stackrel{\triangle}{=} K_2 \text{- correction factor}} \quad \underbrace{\phantom{A_2 \hat{X}_1}}_{\text{error (innovation)}}$

$\Rightarrow \hat{X}_2$ and P_2^{-1} are obtained recursively by using \hat{X}_1 and P_1^{-1} at the previous step \Rightarrow The procedure can be generalized:

$$P_i^{-1} = P_{i-1}^{-1} + A_i^T R_i^{-1} A_i$$

$$K_i = P_i A_i^T R_i^{-1}$$

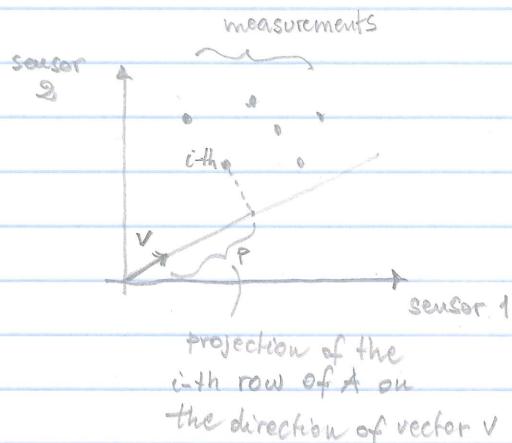
$$\hat{X}_i = X_{i-1} + K_i (Y_i - A_i \hat{X}_{i-1})$$

2) Singular Value Decomposition

$$A = \left[\begin{array}{c|c} \text{e.g., sensors} & n \times m \\ \hline \end{array} \right] \quad \left. \begin{array}{l} \text{e.g., measurements} \\ n \geq m \end{array} \right.$$

$$p = \frac{a_i}{\sqrt{n}} v \Rightarrow \text{For all rows: } \begin{bmatrix} p_1 \\ p_2 \\ \vdots \\ p_n \end{bmatrix} = Av$$

↑
i-th row



v = first singular vector is the vector that maximizes $\|Av\|^2$

⑧

Analogously, one can define:

$v_2 \triangleq$ second singular vector is the vector that maximizes $\|Av\|^2$
conditioned to being perpendicular to v_1

The procedure can be iterated and I can find m orthogonal unit vectors that form a base for the space of rows of A

$$V = [v_1 | v_2 | \dots | v_m]$$

For each v_i , let us call: $\sigma_i \triangleq \|Av_i\|$ - i -th singular value

Now, let us consider the transformation that A performs on the vectors v_i

$u_i \triangleq \frac{1}{\sigma_i} Av_i$ - $n \times 1$ vector \Rightarrow It is the i -th LEFT singular vector of A (it has length = 1)

$$U = [u_1 | u_2 | \dots | u_m]$$

Now note this:

$$\|Av_i\|^2 = v_i^T A^T A v_i = \sigma_i^2 \Rightarrow v_i \text{ is an eigenvector of } A^T A$$

$$\|A^T u_i\|^2 = u_i^T A^T A u_i = \underbrace{\frac{1}{\sigma_i^2} v_i^T A^T A v_i}_{\begin{array}{l} \text{because} \\ \text{of the definition} \\ \text{of } u_i \end{array}} = \frac{1}{\sigma_i^2} \sigma_i^2 v_i^T v_i = \sigma_i^2$$

$\Rightarrow u_i$ is an eigenvector of $A^T A$

Therefore, the following happens:

$$[\sigma_1 u_1 \quad \sigma_2 u_2 \dots \quad \sigma_m u_m] = A V$$

(9)

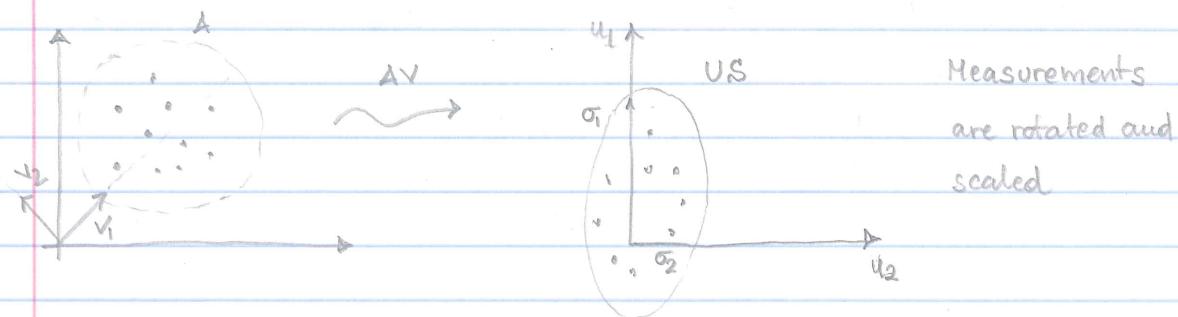
and by defining $S \triangleq \begin{bmatrix} \sigma_1 & & \\ & \sigma_2 & \\ & & \ddots \\ & & & \sigma_m \end{bmatrix}$, one can write:
 diagonal matrix
 $(m \times m)$

$$US = AV \Rightarrow A = USV^T - \text{Singular Value Decomposition of } A$$

Note the construction process \Rightarrow Decomposition is not unique

$\Rightarrow V$ is a set of orthonormal eigenvectors of $A^T A$ (It always exists because $A^T A$ is symmetric and with real entries), i.e.:
 $V^T V = I_m$

Example:



Note: $A^T A$ is an $n \times n$ matrix \Rightarrow It may have n eigenvectors
 Therefore, a typical SVD computes:

$U: n \times n$

$V: m \times m$

$S = \begin{bmatrix} \sigma_1 & & \\ & \sigma_2 & \\ & & \ddots \\ & & & \sigma_m \end{bmatrix}$: $n \times m$ with only up to m nonnegative singular values

(10)

Let us return to the case of A being a matrix of measurements:

$$A = USV^T$$



$n \times m$
↑
temporal information
information

U is $n \times n \Rightarrow$ It is a collection of n temporal "modes"

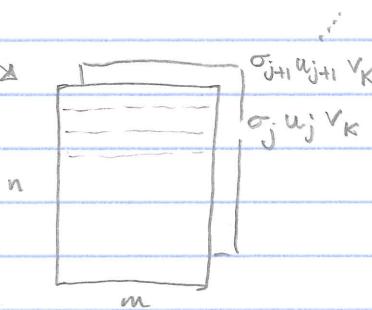
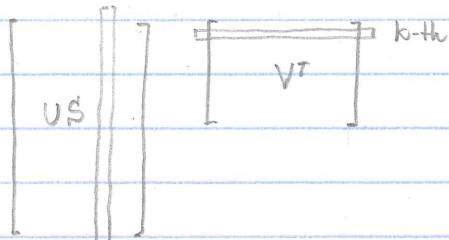
V is $m \times m \Rightarrow$ It is a collection of m spatial "modes"

It is called
"matrix of scores"

US is $n \times m \Rightarrow$ Only m temporal modes are retained and each one is weighted by the correspondent singular value

According to this interpretation, we are saying that the k -th right

singular vector contributes to the measurements collected by every sensor but, at each time point, the contribution is different across temporal modes



Note: we are not adding or removing information. We are simply rearranging information via linear operations

What is the advantage of doing this?

Data compression: Since σ_i are ordered from the largest to the smallest, one can approximate A with the summation of just a few elements in the stack of matrices $\sigma_j u_j v_k$

A criterion to stop: Choose h such that $\sum_{i=1}^h \sigma_i / \sum_{i=1}^m \sigma_i$ is above a

given threshold and use the first h singular vectors \Rightarrow That ratio indicates the probability of the first h components accounting for the energy (or variance) of A

Data separation: Different spatial modes may correspond to different players (e.g., sources) \Rightarrow Data points mainly contributed by one mode are likely related to that source

Noise cancellation: The first singular value (if mean is not removed from the data) and the smallest singular values may be associated with non-physiologically-meaningfull modes \Rightarrow Removal allows to increase the SNR of the data points of interest

